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Assessing the suitability of the minimum capture size and protection regimes in the

gooseneck barnacle shellfishery

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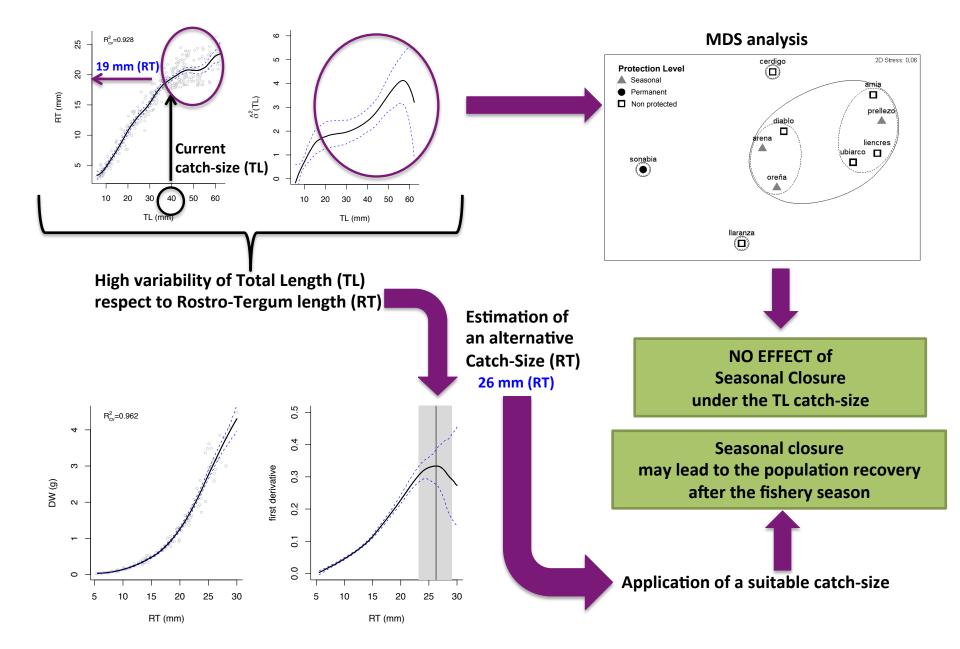
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Minimum capture size

Protection regimes



- Assessing the suitability of the minimum capture size and protection regimes in the
 gooseneck barnacle shellfishery
- 3
- 4 Abstract
- 5

6 The suitability of a total-length-based, minimum capture-size and different protection regimes 7 was investigated for the gooseneck barnacle Pollicipes pollicipes shellfishery in N Spain. For 8 this analysis, individuals that were collected from 10 sites under different fishery protection 9 regimes (permanently open, seasonally closed, and permanently closed) were used. First, we 10 applied a non-parametric regression model to explore the relationship between the capitulum 11 Rostro-Tergum (RT) size and the Total Length (TL). Important heteroskedastic disturbances 12 were detected for this relationship, demonstrating a high variability of TL with respect to RT. 13 This result substantiates the unsuitability of a TL-based minimum size by means of a 14 mathematical model. Due to these disturbances, an alternative growth-based minimum capture 15 size of 26.3 mm RT (23 mm RC) was estimated using the first derivative of a Kernel-based non-16 parametric regression model for the relationship between RT and dry weight. For this purpose, 17 data from the permanently protected area were used to avoid bias due to the fishery. Second, the 18 size-frequency distribution similarity was computed using a MDS analysis for the studied sites 19 to evaluate the effectiveness of the protection regimes. The results of this analysis indicated a 20 positive effect of the permanent protection, while the effect of the seasonal closure was not 21 detected. This result needs to be interpreted with caution because the current harvesting based 22 on a potentially unsuitable minimum capture size may dampen the efficacy of the seasonal 23 protection regime.

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- 25

26

1. Introduction

31	The primary geographical distribution area of the gooseneck barnacle Pollicipes pollicipes
32	ranges from the northwestern coast of France (Brittany) to the northwestern coast of Africa
33	(Senegal) and the Mediterranean (Algeria) (Cruz and Araujo, 1999; Barnes, 1996). This species
34	constitutes the most economically important shellfishery resource on the intertidal rocky shores
35	of Portugal and Spain (Cunha and Weber, 2001, Sousa et al., 2013). This species is highly
36	prized as food (> 50 Euros Kg ⁻¹ (Jacinto et al., 2010)) and heavily exploited by professional and
37	recreational fishery. In recent years, this species has attracted increased harvesting pressure due
38	to its high market value (Sousa et al., 2013; Stewart et al., 2014), the decline of other coastal
39	fisheries has urged barnacle exploitation as a supplement to fishing activity (Bald and Borja,
40	2012), and the European economic crisis. In many regions of the Iberian Peninsula, this decline
41	has resulted in the overexploitation of the stocks (Borja et al., 2011, Stewart et al., 2014).
42	Likely, the current passive management model of the resource (i.e., non-take zones and legal
43	minimal size of capture) of the northern regions of Spain (Cantabria, Basque Country) means a
44	progressive decline of the resources (Jamieson et al., 1999; Bald and Borja, 2012).
45	
46	However, the assessment of the performance of these management measures is not easy. The
47	fishery capture data are often scarce and lack precise localization information, which leads to
48	not very rigorous estimations of the fishery pressure upon this resource (Sousa et al., 2013). The
49	highest abundances of P. pollicipes are located in the lower intertidal zone (Cruz, 2010; Pavón,
50	2003) of significantly energetic shores that are exposed to dominant swells, which are
51	frequently related to high slopes and the presence of caves and crevices (Barnes, 1996; Cruz,
52	2000; et al., 2006: Borja et al, 2006b). Consequently, the poor capture data and the physical
53	factors determining the distribution of the genus Pollicipes and explaining the difficulty of
54	sampling (Bernard, 1988; Parada et al., 2012) contribute to the lack of large-scale population

55 assessment studies and adequate evaluations of the performance of the management measures. 56 As an alternative, the territorial use rights for fishing (TURF), i.e., an area-based management 57 program that assigns a specific area to an individual, group or community, has proven to be an 58 effective approach for the small-scale management of *P. pollicipes* fisheries in NW Spain 59 (Molares and Freire, 2003). The TURF programs grant exclusive fishing access to these 60 communities while giving them management responsibilities, including the development of 61 annual management plans and the maintenance of appropriate controls of fishing mortality 62 (Young, 2013).

63

64 Regardless of the management model, the measures that are oriented toward the achievement of 65 a sustainable exploitation of the resources in N Spain commonly include the minimum size of 66 capture, protected areas (e.g., seasonal or permanent closures) and individual quotas (e.g., 67 Parada et al., 2012; Sousa et al., 2013). In recent years, to react to barnacle population decline in 68 some regions, the establishment of a minimum capture size has received special attention from 69 managers and researchers, particularly regarding the adequacy of the part of the barnacle that is 70 measured. The commonly used capture size in N Spain based on the total length (TL) may be 71 inappropriate because this measure includes both the hard part (capitulum) and the soft part (the 72 peduncle) of the barnacle. The latter has, a priori, an importantly variable typology, with both 73 barnacles with elongated and with smooth peduncles being able to have the same capitulum size 74 (Parada et al., 2012). This finding may lead to heteroskedastic disturbances in the relationship 75 between TL and the capitulum length, i.e., an important variance in TL with respect to the 76 capitulum length as individuals grow in size. This variability in TL depends on environmental 77 factors, such as hydrodynamic patterns, degree of immersion, availability of food and 78 intraspecific competition (Hoffman, 1988, 1989; Page, 1986; Lewis & Chia, 1981). 79 80 Several authors have regionally determined that the rostro-tergum (RT) (Pavon, 2003) or

81 rostrum-carina (RC) (Cruz, 1993) lengths are the best biometrical variables to explain the

82	growth of this species. Consequently, these two measures and others also based on the
83	capitulum have been considered more adequate for the establishment of the minimum size of
84	capture in NW Spain and SW Portugal (Parada et al., 2012; Sousa et al., 2013). Sestelo and
85	Roca-Pardiñas (2011) recently investigated the minimum capture size for this species using a
86	non-parametric model that analyzes the length-weight relationship and its derivatives. This
87	author suggested an RC length-based capture size that ensures the maximum yield in weight
88	from the fishery. To our knowledge, no information has been published on the impact of the
89	change in the minimum capture size from the TL-based measure to a capitulum-based measure.
90	Despite considering the alternative minimum size based on the capitulum length, the assumed
91	lower suitability of the TL measure compared to the RC or RT has not been properly
92	investigated in terms of the variability of the TL with respect to the growth of the species, i.e.,
93	in terms of the heteroskedasticity in TL-RT or TL-RC relationships. However, along the
94	northern coast of Spain (Gulf of Biscay), the minimum capture size is still based on the TL.
95	
96	Regarding the effectiveness of closure regimes in enhancing population stocks, Cruz (2000) and
96 97	Regarding the effectiveness of closure regimes in enhancing population stocks, Cruz (2000) and Sousa et al. (2013) did not find significant effects on the density between areas with different
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109 Kg m⁻²) compared to that in unprotected areas (\sim 1.5 Kg m⁻²). The failure of the temporal

110 closures might be in part linked to the current minimum size in N Spain.

111

112 The extraction of *P. pollicipes* in N Spain is regulated by regional-scale management models 113 largely based in a legal capture size of 40 mm of total length and different types of closure 114 regimes. The purpose of this study was to investigate the effectiveness or suitability of these 115 management measures, considering that the failure of the temporal closures may be in part 116 associated with the potential unsuitability of the current minimum size. For this purpose, the 117 coastline of Cantabria (Gulf of Biscay) of 215 km was selected as a case study due to (i) the fact 118 that the management model in this region permits the analysis of three types of protection 119 regimes (i.e., permanently open, seasonally closed, and permanently closed) and (ii) the lack of 120 previous assessments of the efficacy of these management measures. The conditional variance 121 of the total length with respect to the growth was analyzed (i) to properly assess the suitability 122 of the capture size based on a measure including the soft part of the barnacle and, if unsuitable, 123 (ii) to estimate an alternative capture size using a nonparametric regression model for the 124 capitulum length-weight relationship. The results may confirm with a mathematical model the 125 suitability of the capitulum length-based minimum size measures that are already implemented 126 in other regions. The size-frequency distribution similarity between zones with different 127 protection regimes was analyzed to evaluate their effectiveness. The results of this latter 128 analysis were interpreted according to the differences between the original minimum capture 129 size and the alternative size that is proposed in this study. 130

- 131
- 132 **2. Material and Methods**
- 133

134 2.1. Study area

136 The total area of study covered 215 km of coast in the Cantabria region (Figure 1). Due to the 137 lack of proper habitat suitability data (i.e., mapping), the suitable habitat of gooseneck barnacle 138 was identified using a compilation of information coming from cartographic data, professional 139 shellfishers and technical personnel of the Regional Fishing Directorate. A total of 10 coastal 140 areas of different lengths covering 60 km of coastline were considered common shellfishing 141 zones. These areas were managed under different protection levels according to their fishery 142 closure regime: (1) Llaranza, Ubiarco, Liencres, Arnia, Diablo and Cerdigo were opened to the 143 fishery throughout the year; (2) Arena, Prellezo and Oreña were seasonally protected, being 144 closed to the fishery from 1st May to 1st October; and (3) Sonabia was permanently protected 145 and closed to the fishery throughout the year (Figure 1). A sampling site that was representative 146 of the abundance of each fishing zone was selected based on the compiled information from 147 shellfishermen and technical personnel. The criteria for this selection were to have similar 148 accessibility and wave exposure for every site.

149

150 **Figure 1**

151

152 2.2. Sampling and laboratory procedures

153

154 Field surveys were carried out during the spring low tides in June 2005, except at Diablo, which 155 was surveyed in October 2005. At each of the 10 sites, 5 intertidal transects were established 156 along 200-300 m of coastline, except in Cerdigo and Diablo (4 transects) and in Llaranza (3 157 transects). At each transect, three different intertidal levels were studied: higher littoral (H), 158 medium littoral (M) and lower littoral (L). The shallow subtidal zone was not considered 159 because this species is mainly distributed and professionally harvested in the intertidal zone 160 (Borja et al., 2006a). The individuals were collected by scraping the rock surface within $50 \times$ 161 50-cm sampling units. In the laboratory, several biometric variables were measured: (i) Rostro-162 tergum length (RT) (Pavón, 2003), also known as the maximal capitulum height (MCH) by

163	Cruz (1993) or the capitulum length by Bernard (1988); (ii) rostro-carina length (RC) following
164	Cruz (1993); and (iii) total length (TL) (Hoffman, 1984) (Figure 2). The dry weight (DW) was
165	obtained by drying organisms at 60 °C for 48 h following Cruz (1993). Encrusting and boring
166	animals, mainly located on the capitulum plates, were removed before weighing.
167	
168	Figure 2
169	
170	2.3. Data analysis
171	
172	2.3.1. Evaluation of the total length-based minimum capture size
173	
174	The suitability of the TL for use as a minimum capture size measure was evaluated by analyzing
175	the conditional variance $\hat{\sigma}^2$ of TL, which includes both the soft part of the peduncle and the
176	hard part of the capitulum, with respect to the capitulum length, using the statistical software
177	environment R (R Development Core Team, 2009). $\hat{\sigma}^2$ values that were larger than 1.5 were
178	considered significant in terms of TL variability with respect to RT. The RT size was selected as
179	the capitulum length measure based on a locality criterion, as it was determined by Pavon
180	(2003) as the best biometrical variable to explain the growth of this species in the neighboring
181	coast of Asturias (N Spain). Cruz (1993) found the RC size to be more adequate for this purpose
182	in SW Portugal. In this study, the RC-RT relationship was also first determined to compare our
183	and other results based on the RT size with studies based on the RC size (Figure 3).
184	
185	The analysis of RT-TL relationship was conducted using a random subsample of the whole
186	dataset, including all sites and tidal levels. A Kernel-based non-parametric regression model
187	was used for this analysis. The model equation was of the type
188	
189	$RT = m(TL) + \sigma(TL) \varepsilon \qquad (1)$

191 where the error variable ε is independent of the covariate TL with $E(\varepsilon) = 0$ and

- 192 $Var(\varepsilon) = 1, m(TL) = E(RT|TL)$ is the unknown regression function, and $\sigma^2(TL) =$
- 193 *Var(RT|TL)* is the conditional variance function representing heteroskedasticity.
- 194
- 195 The estimation of the above model is based on the use of local linear kernel smoothers (Wand

and Jones, 1995). Let $\{(TL_i, RT_i)\}_{i=1}^n$ be an independent random sample, the kernel estimator

197

$$\widehat{m}(tl) = \psi(tl, \{(TL_i, RT_i)\}_{i=1}^n, h)$$

198

199 at a location tl is defined as $\hat{m}(tl) = \hat{\alpha}_0(tl)$, where $\hat{\alpha}_0(tl)$ is the first position of the vector

200 $(\hat{\alpha}_0(tl), \hat{\alpha}_1(tl))$, which is the minimizer of

201

$$\sum_{i=1}^{n} \{RT_i - \alpha_0(tl) - \alpha_1(tl)(TL_i - tl)\}^2 K_h(TL_i - tl),$$

202

203 where
$$K_h(\cdot) = K(\cdot/h)/h$$
, $K(\cdot)$ denotes a kernel function (a symmetric density), and $h > 0$ is

204 the smoothing parameter (or bandwidth). Additionally, the residual sum of square approach was

205 used to estimate the conditional variance as

$$\hat{\sigma}^2(tl) = \psi(tl, \{(TL_i, R_i)\}_{i=1}^n, h_R)$$

206 with $R_i = (RT_i - \hat{m}(TL_i))^2$.

- 208 Finally, to determine the model's adjustment, the coefficient of determination was used, as
- 209 calculated by means of cross-validation (Stone, 1977). This calculation is

$$R_{cv}^{2} = 1 - \frac{\sum_{i=1}^{n} (Y_{i} - \hat{Y}_{i}^{(-i)})^{2}}{\sum_{i=1}^{n} (Y_{i} - \bar{Y})^{2}}$$

210 where $\hat{Y}_i^{(-i)}$ indicates the estimates of Y_i leaving out the *i*-th element of the sample as obtained 211 by fitting the corresponding model and $\bar{Y} = n^{-1} \sum_{i=1}^{n} Y_i$.

212

213

214

- 215 2.3.2. Estimation of the minimum size of capture
- 216

The conditional variance ($\hat{\sigma}^2$) analysis of TL with respect to RT showed values of $\hat{\sigma}^2$ that were 217 218 larger than 1.5 beyond TL=15 mm. This result confirmed the current TL-based measure as 219 unsuitable for setting the minimum size of capture. Consequently, an alternative catch size was 220 estimated based on the growth-based RT length measure. Prior to this estimation, the RT size 221 corresponding to the current minimum size of capture in Cantabria (i.e., TL=40 mm) was 222 estimated using the regression model of the equation (1). This estimation permitted the current 223 TL size in RT to be comparable to the alternative catch size that we estimated as follows. 224 225 For individuals that collected from the permanently protected zone of Sonabia, a minimum 226 capture size was estimated using the methodology of Sestelo and Roca-Pardiñas (2011) by 227 means of the RT-DW length-weight relationship. The fact that the analysis was conducted using 228 data from a permanent protected zone ensured that results were not affected by the fishery 229 pressure. 230 231 The following nonparametric model was applied to study the RT-DW relationship 232 $DW = m^*(RT) + \sigma^*(RT) \varepsilon^* (2)$ 233 234

235	where m^* is the unknown regression function, ${\sigma^*}^2$ is the conditional variance function, and ε^* is
236	the error independent of the covariate RT. The estimation of this model was obtained
237	analogously to the model in (1).
238	
239	The first derivative of m^* was calculated to determine an RT-based new suitable size of capture
240	for this species. The ideal size, named rt_0 , was used as the maximizer of the first derivative of
241	m^* . This point could be define as
242	
243	$rt_0 = arg \max_{rt} m^{*(1)}(rt).$ (3)
244	
245	Beyond this point, the increase in weight per unit of size decreases. Thus, this size ensures that
246	individuals thatsmaller than this size had not yet attained the maximum yield in weight. See a
247	full description of this methodology in Sestelo and Roca-Pardiñas (2011). This RT-DW
248	relationship was also analyzed using the classic allometric model (Huxley, 1924) only with the
249	purpose of comparing the results of both of the models (see Figure A1 in the Appendix). Note
250	that the first derivative of the allometric model cannot show a maximum due to its continuously
251	increasing nature.
252	
253	2.3.3. Size frequency distributions
254	
255	Size frequency distributions with discrete size classes of 5-mm RT intervals were obtained for
256	each site. The population class distribution was also analyzed at each tidal level for the entire
257	dataset. For the descriptive analyses of recruitment patterns and exploitable stocks, two barnacle
258	sizes were selected: RT<5 mm and RT>20 mm. These sizes were assumed to correspond to the
259	previous year recruiters and to the minimum size of capture that is allowed in Cantabria,
260	respectively. The capture size is 40 mm TL, with an approximately 20-mm RT length according
261	to the value that was obtained in the present study (see Figure 4a).

263 2.3.4. Evaluation of the protection regimes

-0.	
265	Using the composition of the size frequency distributions at each site as the input data, a
266	similarity matrix was constructed to perform a multidimensional scaling analysis (MDS) (using
267	Manhattan distances) with PRIMER 6.0 software (Primer-e Ltd, 2006). The results of the
268	ordination analyses were used to identify similarities and differences in the population size
269	structure among sites with different protection levels. Similar sites were then grouped using the
270	results that were provided by a single linkage cluster analysis of the same data. The criterion to
271	evaluate the effectiveness of the different protection levels was the following: the sites that were
272	grouped in the same group of non-protected sites were considered to have an ineffective
273	protection regime, while sites that were grouped in a different group to that of non-protected
274	sites were considered to have an effective protection regime. This type of analysis was
275	previously applied, with similar purposes, in SW Portugal by Cruz et al. (2010) and Sousa et al.
276	(2013) for <i>P. pollicipes</i> .
277	
277 278	3. Results
	3. Results
278	3. Results3.1. Evaluation of the total length–based minimum capture size
278 279	
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 278 279 280 281 282 283 284 	 3.1. Evaluation of the total length–based minimum capture size The RT-TL relationship was explained using a Kernel-based nonparametric regression model (Figure 4a). Once the model was applied, it was possible to clearly observe the nonparametric increasing trend of the estimated mean (Figure 4a) and the heteroskedastic disturbances (Figure 4a)
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 278 279 280 281 282 283 284 285 286 	3.1. Evaluation of the total length–based minimum capture size The RT-TL relationship was explained using a Kernel-based nonparametric regression model (Figure 4a). Once the model was applied, it was possible to clearly observe the nonparametric increasing trend of the estimated mean (Figure 4a) and the heteroskedastic disturbances (Figure 4b). The conditional variance $\hat{\sigma}^2$ with respect to RT significantly increases beyond a TL size of 15 mm (Mean=1.5, CI=(0.9, 1.9)) to display a maximum at TL=58 mm (Mean=4.0, CI=(1.1,

289	obtaining the r	naximum yield	in weight from	the fishery becau	se a great percentage of
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290 individuals could correspond to harvestable sizes with an alternative capitulum-based minimum

- size (see Figure 4a: data with TL<40 mm and RT>19.4 mm) and (ii) harvest individuals clearly
- smaller than the correspondent capitulum-based minimum size (see Figure 4a: data with TL>40
- 293 mm and RT<19 mm). Note that RT=19.4 mm was obtained in section 3.2.

294

295 Figure 4

296 3.2. Minimum size of capture

Based on the nonparametric model that was used to analyze the RT-TL relationship (Figure 4a)

and being cautious to interpret the result due to this model's heteroskedastic disturbances, it is

299 possible to suggest a capture size of 19.4 mm (19.0, 19.8) in RT length corresponding to the

300 current 40-mm TL legal capture size. This result needs to be taken with caution due to the

301 observed variability of TL with respect to RT. Using the RC-RT relationship that was obtained

in the present study, the corresponding size of the RC was 16.9 (16.8, 17.0) (Figure 3).

303

Figure 5

305

306 Alternatively and due to the significant heteroskedasticity in the RT-TL relationship, with an 307 increase in the variance as individuals grow in size (Figure 4b), an alternative minimum capture 308 size based on the RT-DW length-weight relationship was estimated. Thus, for individuals that 309 were collected from the permanently protected area of Sonabia, the Kernel-based nonparametric 310 model showed that the regression curve was an increasing function (Figure 5a) that was very 311 similar but displayed some differences in the last part of the curve from that obtained by the 312 allometric model (Figure A1). The first derivative of the initial curve displayed an increasing 313 monotonous function in the first part of the curve and a maximum at a specific size, after which 314 the curve began to decrease (Figure 5b). The size at which the maximum value was obtained for 315 the derivative was 26.3 mm (23.2, 29.1) in RT and was considered the minimum suitable catch

size. The RC-RT regression model (Figure 3) was applied to obtain a minimum suitable catch
size of 23 mm (22.8, 23.2) in RC length.

318

319

320 3.3. Population size structure

321

322 Size frequency distribution plots for the 10 studied sites and for the three tidal levels are shown 323 in Figure 6. Oreña, Llaranza, Arena and Sonabia stood out appreciably from the rest of the sites 324 with a marked main mode at RT sizes <5 mm. Apart from the <5-mm sizes, the next mode 325 corresponded to the 15-20-mm range in all sites, except for in Prellezo, Diablo and Cerdigo, 326 where the main size mode corresponded to the ranges between 5 and 15 mm, even exceeding 327 the <5 mm range. The relative percentages of individuals of exploitable sizes (>20 mm) were 328 consistently higher at the permanently protected zone of Sonabia (19%). Arnia and Cerdigo, 329 which were opened to the fishery, also showed important percentages of commercial-size 330 individuals (12-15%). The average size structure plots by tidal level (Figure 6b) showed a main 331 mode corresponding to individuals of <5 mm, especially marked at high and medium levels, and 332 a secondary mode in the 15-20-mm range. From that size range onwards, there was a rapid 333 decrease in the commercial sizes (>20 mm), which are mainly concentrated at the medium 334 levels and especially at the lower tidal levels. 335 336 Figure 6 337 338 3.4. Evaluation of the protection regimes 339 340 The ordination analysis that was carried out by Multi Dimensional Scaling (MDS) using size 341 frequency distributions showed that the permanently protected site of Sonabia clearly separated

342 from the remainder sites (Figure 7). A Similarity Percentage analysis, SIMPER (Clarke 1993),

343 was performed to compare the size frequency distributions between Sonabia and the rest of the 344 sites and suggested that the main cause of these differences was the higher percentage of large 345 sizes found in Sonabia. The Manhattan distances that were used to group the sites with similar 346 population size structure (16 and 18) grouped 7 sites without showing any distribution pattern 347 that was associated with seasonal closure or non-protection regimes. Cerdigo and Llaranza, both 348 non-protected, were individually grouped apart, being located closer to the 7 grouped sites than 349 to Sonabia. 350 351 Figure 7 352 353 4. Discussion 354 355 The results of this study suggest that the two most commonly used management measures 356 within the commercial goose barnacle (*Pollicipes pollicipes*) fishery in N Spain, with a 357 minimum capture size of 40 mm TL and seasonally closed areas, are not entirely suitable, at 358 least together, for a sustainable harvesting of the resource. These results are discussed below in 359 the context of the sustainable exploitation of this resource in the studied coast and their 360 transferability to other coastal regions. The result regarding the suitability of temporal closures 361 is interpreted in light of the difference between the current capture size and the alternative size

363 studies, a regression model of the relationship between the capitulum RT and RC biometrical

that were proposed in this study. To discuss our results in comparative terms regarding previous

364 measures was formulated. This model shows that RC length was approximately 90% of the RT

365 length (Figure 3), suggesting transforming the data when these measures are compared. The

366 good fit of the model suggests that both of the biometrical variables could be appropriate and

asily comparable.

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369 The peduncle of the barnacle produced heteroskedasticity in the RT-TL relationship, displaying 370 an important variance in TL with respect to the growth-based measure RT as individuals grow 371 in size (Figure 4). This result substantiates with a mathematical model the unsuitability of the 372 TL-based minimum capture size, which constitutes the most common reference measure for the 373 legal minimum size of capture in N Spain, and analytically supports the capitulum measure-374 based (i.e., growth-based) catch size that is already adopted in several regions of the Iberian 375 Peninsula (Parada et al., 2012; Sousa et al., 2013). Consequently, a new minimum suitable size 376 was estimated following the methodology of Sestelo and Roca-Pardiñas (2011) and based on the 377 RT capitulum length. The first derivative of the nonparametric model showed a suitable 378 minimum capture size of 26.3 mm RT, which corresponds to 23 mm RC using the RT vs RC 379 relationship (Figure 3). The size that was obtained with this methodology ensures the maximum 380 yield in weight from the fishery. This size is clearly larger than the current catch size in N Spain 381 (40 mm TL) considering that the parametric regression model that was obtained in the present 382 study for the RT-TL relationship (Figure 4a) suggested a minimum size of capture of 19.4 mm 383 RT (or 16.9 mm RC) corresponding to the current minimum size of 40 mm TL. 384

385 The results that were obtained in this study agree with the size limit that was recently proposed 386 by Jacinto et al. (2011) in Portugal (50% of the total harvested biomass comprising individuals 387 > 22 mm RC). Sestelo and Roca-Pardiñas (2011) recently investigated the minimum capture 388 size for this species in NW Spain (Galicia) using the same non-parametric model as that applied 389 in this study. Their study, based on RC, estimated a suitable minimum capture size of 21.5 mm, 390 also ensuring that any barnacle under this size has not yet attained its maximum yield in weight. 391 Fishing directorates in different regions of Spain and Portugal have already used the RC length 392 to establish the minimum size of capture. For instance, in Portugal, the capture size is between 393 20 mm and 23 mm in RC length (Sousa et al., 2013), and in Galicia the way to measure the 394 minimum size of capture was changed from the TL (40 mm) to the capitulum-based (CB) length

- (15 mm) (Parada et al., 2012). Using the CB vs RC regression model of Parada (2013), the
 minimum capture size for Galicia should be 18.3 mm RC.
- 397

398 To a greater or lesser extent, these sizes are above the 12.5-mm RC (14.26 mm RT) sexual 399 maturation size (Cruz, 2000) and consequently may permit the production of a minimum of 1-2 400 broods before the designated size is reached (Molares et al., 1994; Pavón, 2003; Sestelo and 401 Roca-Pardiñas, 2011). However, the current minimum capture size in Cantabria (40 mm TL or 402 16.9 mm RC) may be less conservative in terms of the sustainability of the fishery compared to 403 the rest of analyzed capture sizes. The important difference that was observed in the minimum 404 size between the current capture size in N Spain or the capture size that was established in 405 Galicia (18.3 mm RC) and that obtained in our study (23 mm RC) is explained by the highly 406 conservative approach of Sestelo and Roca-Pardiñas (2011). This method to estimate the 407 minimum size of capture not only ensures at least a brood before the designated size is reached 408 but also attains, as discussed, the maximum yield in weight from the fishery. It is also worth 409 mentioning that previous studies estimated the minimum suitable size using individuals from 410 open fishing areas, while our analysis was conducted using data from a permanently protected 411 zone. This analysis ensured that the minimum capture size was estimated using data that were 412 not biased by the fishing pressure. This analysis may also contribute to the slightly larger 413 capture size that was obtained in this study compared to that of Sestelo and Roca-Pardiñas 414 (2011). The transferability of this result to other coastal areas to support regional or zone-based 415 management models could be adequate for the sustainability of the resource in highly exploited 416 areas considering that our result is the most conservative. However, for this purpose, further 417 analysis is suggested, including more data (i.e., monthly data). 418

In addition to the methodology that was used and the environmental differences between study
areas, the lack of a standardized measure for the RC length could also explain, in part, the
differences between the above-described minimum capture sizes. The comparison between the

422 minimum sizes of different studies was performed in terms of the RC length. For this purpose, 423 the given size measure, CB or RT, was transformed to RC using the regression models from 424 Parada et al. (2012: 2013) and that obtained in the present study (Figure 3), respectively. 425 However, the RC length that was measured by Parada et al. (2012, Figure 4) and Sestelo and 426 Roca-Pardiñas (2011, Figure 2) is slightly different from that measured in this study (Figure 2) 427 following Cruz (1993, Figure 2). We chose to measure the RC as did Cruz (1993) because she 428 demonstrated that it was the biometrical measure that best represented the linear growth of P. 429 *pollicipes* and is used to establish the minimum capture size in Portugal.

430

431 The permanent protection regime had a positive effect on the population structure in terms of 432 the abundance of large-size individuals (Figure 6). The MDS analysis demonstrated that the 433 permanently protected site of Sonabia clearly separated from the non-protected sites group 434 (Figure 7). This result is in agreement with those of Borja et al. (2006b), who observed that the 435 density and biomass were 5 times higher in permanently unexploited zones than in unprotected 436 areas in the neighboring coast of Basque Country. However, Cruz (2000) and recently Sousa et 437 al. (2013) in a similar analysis did not find any positive significant effect on the percentage of 438 cover or the density of large individuals in no-take areas of central and SW Portugal. Sousa et 439 al. (2013) suggested that the absence of this effect could be observed because the restrictions of 440 exploitation are frequently not respected. However, seasonally protected sites (from May to 441 October) did not show any significant differences in the size frequency distribution patterns 442 compared to those of non-protected sites, i.e., the MDS analysis mainly grouped seasonal 443 closure sites and non-protected sites in the same group (Figure 7). Bald et al. (2006) obtained 444 similar results in the simulations that were conducted to reproduce different closure scenarios 445 using a system dynamic model. These authors considered that 5-7 months of temporal closures 446 are not sufficient to permit a total recovery of the resource after the capture season and 447 suggested an annual alternate exploitation of the fishing zones as the best management decision. 448

449 The descriptive analysis of the size frequency distributions (Figure 5) completed the 450 understanding of the similarities between the sites as observed in the MDS analysis (Figure 6) 451 and confirmed the absence of a positive effect as associated with temporal closures. Sonabia and 452 Cerdigo presented the highest abundance of 20-mm-size individuals and the unique sites that 453 present individuals greater than 25 mm. The case of Sonabia, which was clearly separated from 454 the other sites in the MDS analysis, can be easily explained by the permanent closure of the 455 zone, as mentioned above, while the case of Cerdigo is more probably linked to both of the low 456 coverages that were observed in this coastal zone (Cantabria, 2005), which shows little interest 457 for professional shell-fishermen, and the presence of a high number of deep crevices and a rock 458 islet which are not very accessible to poachers. The analysis of size frequency distribution 459 patterns as conducted by the tidal level showed that larger sizes are found at the lower intertidal 460 level, which is in accordance with previous research in Portugal (Cruz, 2000: Sousa et al., 2013) 461 and N Spain (Pavón, 2003: Borja et al., 2006b). The difficulties that are associated with fishing 462 at the lower tidal levels (Molares and Freire, 2003) might limit the exploitation intensity, and 463 the abundance of large-size barnacles could remain high at this tidal level. Although these 464 results regarding the tidal levels did not contribute to the goals of this study, the observed 465 pattern demonstrates that the population of this coastal region is behaving as do the neighboring 466 zones under fishing pressure.

467

468 At first glance, these results suggest temporal closure as a non-suitable protection regime for 469 assuring the protection and recovery of a population. However, the desirable effect of the 470 temporal closures could be shaded by the fact that fishery is conducted based on a potentially 471 unsuitable catch-size. The results should be interpreted with caution considering that they were 472 obtained under an unsuitable TL-based capture size. The fact that an important percentage of the 473 individuals that extracted using the TL size could be less than a suitable minimum size may lead 474 to important changes in the size-structure of the population and consequently obscure the real 475 effect of the protection in the MDS analysis. The population size distribution would be

476 appreciably different under the alternative capitulum-based minimum size, and in that case, the 477 6-month seasonal closure could have a positive effect, leading to the recovery the population 478 after the capture season. Nevertheless, this seasonal protection regime may not be adequate for 479 highly overexploited populations. A longer fishery closure may be more suitable in these 480 situations requiring a total recovery. Considering the growth rate of this species (Gutiérrez-481 Cobo, 2012; Cruz et al. 2000) for overexploited areas, a minimum closure period of 2.5 years 482 may be considered, which is the time that is required for goose barnacles to achieve a 483 commercial size in Cantabria. 484

485 **5.** Conclusions

486

487 In summary, the analysis in this paper demonstrated that the minimum capture size based on the 488 growth-based capitulum length measure is more adequate than that based on the TL measure. 489 We suggest a catch size of 23 mm (RC) or 26 mm (RT) in N Spain that is larger than the current 490 catch size that is established in this coast's regions and similar to that proposed in Portugal. The 491 establishment of this size using the first derivative of the length-weight relationship may ensure 492 the maximum yield in weight from the fishery. A seasonal closure of 6 months may not be 493 effective for assuring the total recovery of a highly exploited population after the fishing season, 494 at least if it is not linked to the establishment of a more suitable catch size than the current catch 495 size. Thus, a 6-month seasonal closure should be locally tested in the field under the new 496 minimum size. We also propose a temporal closure of 2.5 years for highly exploited populations 497 to completely recover. Although the analysis applied in this study may be applicable to other 498 coastal areas, the generalizability and transferability of these results to support management 499 decisions at a larger spatial scale are subjected to further research, including monthly data and 500 field experiments, to test the suitability of the obtained results. 501

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505	Appendix
506	(Figure A1)
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523	References
524	
525	Bald, J., Borja, A., Muxika, I., 2006. A system dynamics model for the management of the
526	gooseneck barnacle (Pollicipes pollicipes) in the marine reserve of Gaztelugatxe (Northern
527	Spain). Ecol. Model. 194, 306–315.
528	

529	Bald, J., Borja, A., 2012. The gooseneck barnacle (Pollicipes pollicipes) in the Basque Country
530	(Northern Spain): facing the future of is management. Revista de Investigacion Marina 19 (6),
531	467-468.
532	
533	Bernard, F.R., 1988. Potential fishery for the gooseneck barnacle Pollicipes polymerus
534	(Sowerby, 1833) in British Columbia. Fish. Res. 6, 287–298.
535	
536	
537	Barnes, M., 1996. Pedunculate cirripedes of the genus Pollicipes. Oceanogr. Mar. Biol. Annu.
538	Rev. 34, 303–394. Maritime Engineering and Technology – Guedes Soares et al.
539	
540	Borja, A., Muxika, I., Bald, J., 2006a. Protection of the goose barnacle Pollicipes pollicipes,
541	Gmelin, 1790 population: the Gaztelugatxe Marine Reserve (Basque Country, northern Spain).
542	Sci. Mar. 70, 235–242.
543	
544	Borja, A., Liria, P., Muxika, I., Bald, J., 2006b. Relationships between wave exposure and
545	biomass of the goose barnacle (Pollicipes pollicipes, Gmelin, 1790) in the Gaztelugatxe Marine
546	Reserve (Basque Country, northern Spain). ICES J. Mar. Sci. 63, 626-636.
547	
548	Clarke, K.R., 1993. Non-parametric multivariate analyses of changes in community
549	structure. Australian Journal of Ecology 18, 117–143.
550	
551	Cruz, T., 1993. Growth of Pollicipes pollicipes (Gmelin, 1790) (Cirripedia, Lepadomorpha) on
552	the SW coast of Portugal. Crustaceana 65, 151–158.
553	
554	Cruz, T., 2000. Biologia e ecologia do percebe, Pollicipes pollicipes (Gmelin, 1790), no litoral
555	sudoeste português. PhD Thesis. Universidade de Évora, Portugal.

557 558	Cruz, T., Castro, J.J., Hawkins, S.J., 2010. Recruitment, growth and population size structure of
559	Pollicipes pollicipes in SW Portugal. J. Exp. Mar. Biol. Ecol. 392, 200–209.
560	
561	Cunha, I., Weber, M., 2001. Actividad reproductora del percebe Pollicipes pollicipes en el norte
562	de Portugal. En: D.G.d. P. y. A.G.d. Cantabria (Ed.), Libro de resúmenes VII CongresoNacional
563	de Acuicultura. Santander. pp. 173–175.
564	
565	
566	Gutiérrez-Cobo, M.B., 2012. Recruitment, growth and population size structure of Pollicipes
567	pollicipes on the coast of Cantabria, Gulf of Biscay (N Spain). Considerations for fisheries
568	management. Master of Science Thesis. University of Cantabria. Spain.
569	
570	Hoffman, D.L., 1984. Size-frequency distribution patterns of the juvenile stages of the
571	pedunculate barnacle, Pollicipes polymerus Sowerby, 1833 (Cirripedia, Lepadomorpha).
572	Crustaceana 46, 295–299.
573	
574	Hoffman, D. L., 1988. Settlement and growth of the pedunculate barnacle Pollicipes polymerus
575	Sowerby in an intake seawater system at the Scripps Institution of Oceanography, La Jolla,
576	California.Pacific Science 42, 154-159.
577	
578	Hoffman, D.L., 1989. Settlement and recruitment patterns of a pedunculate barnacle, Pollicipes
579	polymerus Sowerby, off La Jolla, California. Journal of Experimental Marine Biology and
580	Ecology 125(2), 83-98.
581	
582	Huxley, J.S., 1924. Constant differential growth-ratios and their significance. Nature 114, 895-
583	896.

585	IH Cantabria.	2005.	Evaluación	de recursos	de interés	marisquero	en el litora	ıl de	Cantabria	v
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- 586 desarrollo de protocolos aplicables a su gestión. Technical report. University of Cantabria.
- 587 Gobierno de Cantabria, pp. 87 (In Spanish).
- 588
- 589 Jamieson, G.S., Lauzier, R., Gillespie, G., 1999. Phase 1 Framework for undertaking an
- 590 ecological assessment for outer coast rocky intertidal zone. Can. Stock Assess. Secretariat Res.
- 591 Doc. 99, 209.
- 592
- 593 Jacinto, D., Cruz, T, Silva, T., Castro, J.J., 2011. Management of the stalked barnacle
- 594 (Pollicipes pollicipes) fishery in the Berlengas Nature Reserve (Portugal): evaluation of bag and
- 595 size limit regulation measures. Scientia Marina 75 (3), 439-445.
- 596
- 597 Jacinto, D., Cruz, T., Silva, T., and Castro, J. J., 2010. Stalked barnacle (*Pollicipes pollicipes*)
- 598 harvesting in the Berlengas Nature Reserve, Portugal: temporal variation and validation of
- 599 logbook data. ICES Journal of Marine Science, 67: 19–25.
- 600
- 601 Lauzier, R. B., 1999. A review of the biology and fisheries of the goose barnacle (Pollicipes
- 602 polymerus Sowerby, 1833). Fisheries and Oceans Canada, Canadian Stock Assessment
- 603 Secretariat Research Document, 99/111. 30 pp.
- 604
- 605 Lewis, C. A., Chia F., 1981. Growth, fecundity, and reproductive biology in the pedunculate
- 606 cirripede *Pollicipes polymerus* at San Juan Island, Washington. Canadian Journal of Zoology 59,
- 607 893-901.
- 608

609	Molares, J., Tilves, F., Quintana, R., Rodriguez, S., Pascual, C., 1994. Gametogenesis of
610	Pollicipes cornucopia (Cirripedia: Scalpellomorpha) in North-west Spain. Marine Biology 120,
611	553-560.
612	
613	Molares, J., Freire J., 2003. Development and perspectives for community-based management
614	of the goose barnacle (Pollicipes pollicipes) fisheries in Galicia (NW Spain). Fish. Res. 65,
615	485–492.
616	
617	Page, H.M., 1986. Differences in population structure and growth rate of the stalked barnacle
618	Pollicipes polymerus between a rocky headland and an o shore oil platform. Marine Ecology
619	Progress Series 29(2), 157-164.
620	
621	Parada, J.M, Outeiral, R., Iglesias, E., Molares, J., 2012. Assessment of goose barnacle
622	(Pollicipes Gmelin, 1789) stocks in management plans: design of a sampling program
623	based on the harvesters' experienceICES J. Mar. Sci. 69(10), 1840 -1849.
624	
625	Parada, J.M., 2013. A correction to "Assessment of goose barnacle (Pollicipes pollicipes
626	Gmelin, 1789) stocks in management plans: design of a sampling program based on the
627	harvesters' experience''. ICES Journal of Marine Science 70 (1), 244.
628	
629	Pavón, C., 2003. Biología y variables poblacionales del percebe, Pollicipes pollicipes (Gmelin,
630	1790) en Asturias. PhD Thesis.Universidad de Oviedo, Spain.
631	
632	R Development Core Team, 2009. R: A Language and Environment for Statistical Computing.
633	R Foundation for Statistical Computing. Vienna, Austria. ISBN 3-900051-07-0.
634	

635	Sestelo, M., Roca-Pardiñas J., 2011. A new approach to estimation of length-weight relationship
636	of Pollicipes pollicipes (Gmelin, 1789) on the Atlantic coast of Galicia (Northwest Spain): some
637	aspects of its biology and management. Journal of Shellfish Research 30, 939-948.
638	
639	Sousa, A., Jacinto, D., Penteado, N., Martins, P., Fernandes, J., Silva, T., Castro, J.J., Cruz, T.,
640	2013. Patterns of distribution and abundance of the stalked barnacle (Pollicipes pollicipes) in
641	the central and southwest coast of continental Portugal, Journal of Sea Research 83, 187-194.
642	
643	Stewart, A., Fragoso, B. D., Clímaco, R., Icely, J. D., 2014. Evaluation of stakeholder
644	perspectives on the management of the stalked barnacles (Pollicipes pollicipes) resource in the
645	Parque Natural do Sudoeste Alentejano e Costa Vicentina, Portugal. Marine Policy 43, 71-79.
646	
647	Young. J., 2013. Catch Shares in Action: Spanish Galicia Goose Barnacle Cofradía System.
648	Environmental Defense Fund, Technical Report, 10 pp.
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651	Figure captions
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653	Figure 1. Location of sampling sites on the coast of Cantabria (N Spain). * Seasonally protected
654	and ** permanently protected zones. Sites without label represent open fishing zones.
655	
656	Figure 2. Longitudinal biometric variables measured in <i>P. pollicipes</i> following Cruz (1993).
657	RT= Rostrum-Tergum length, RC= Rostrum-Carina length, TL= Total length.
658	Figure 3. Relationship between Rostrum-Tergum length (RT) and Rostrum-Carina length (RC)
659	(n=365). Coefficient of determination (R_{cv}^2) calculated by means of cross-validation is
660	presented for the regression model. Confidence intervals (dashed lines) are hardly identifiable
661	due to the good fit of the model.

Figure 4. Regression curve with bootstrap-based 95% confidence intervals (dashed lines) for Total Length (TL) and Rostrum-Tergum length (RT) (n=365) (a) and the conditional variance of TL (b). R_{cv}^2 represents the coefficient of determination for the regression curve calculated by cross-validation.

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Figure 5. Regression curve (a) and first derivative (b) with bootstrap-based 95% confidence

670 intervals (dashed lines) for dry weight (DW) and Rostrum-Tergum length (RT). Solid vertical

- 671 line (b) represents the estimated rt_0 or the size at which the derivative has the maximum
- 672 value. R_{cv}^2 represents the coefficient of determination for the regression curve calculated by
- 673 cross-validation.

674

675 Figure 6. Size frequency distribution plots (% of individuals) of *P. pollicipes*, (a) for each

676 sampling site and (b) for all data at each tidal level. RT: rostro-tergum length.

677

- 678 **Figure 7.** MDS ordination of the studied sites showing *P. pollicipes* fishery protection levels
- 679 (Circle: permanently closed fishery; Triangle: seasonally closed fishery; Square: null protection
- or all year open to the fishery). Manhattan distances are represented by dotted line (distance=16)

681 and solid line (distance=18).

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683 Appendix figure captions

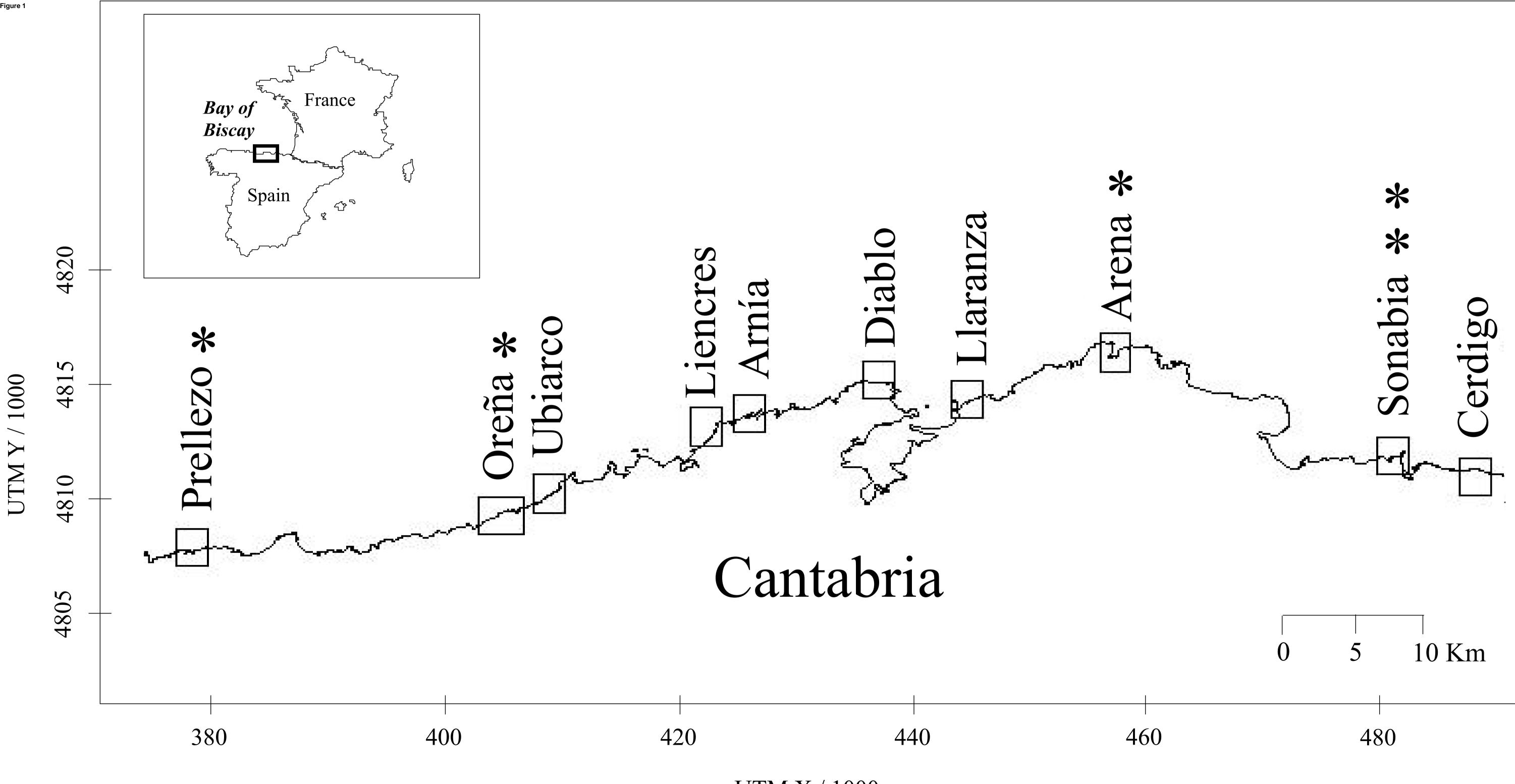
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685 Figure A1. Relationship between Rostro-Tergum length (RT) and Dry Weight (DW) (n=1200).

686 Coefficient of determination (R_{cv}^2) calculated by means of cross-validation is presented for the

- 687 regression model. Confidence intervals (dashed lines) are hardly identifiable due to the good fit
- of the model.





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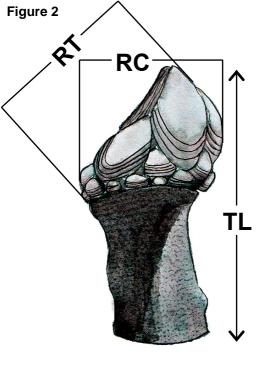


Figure 3

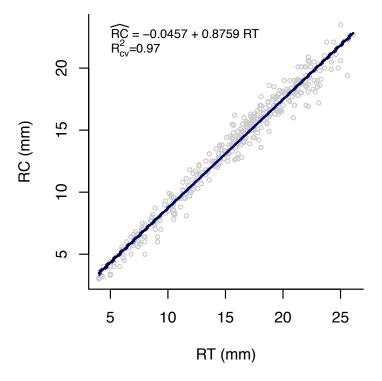


Figure 4

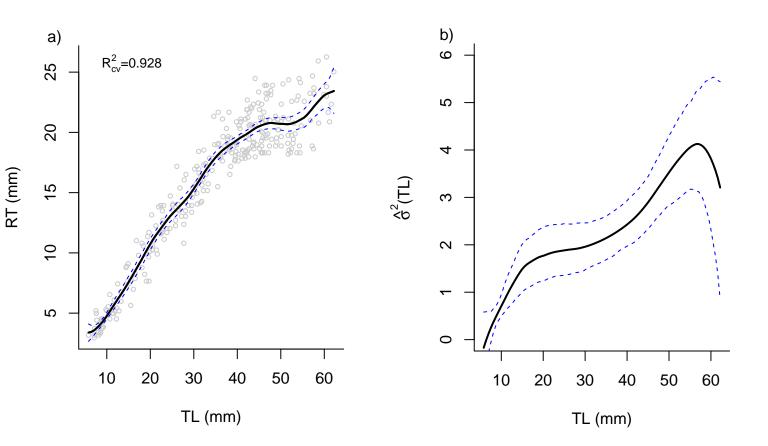
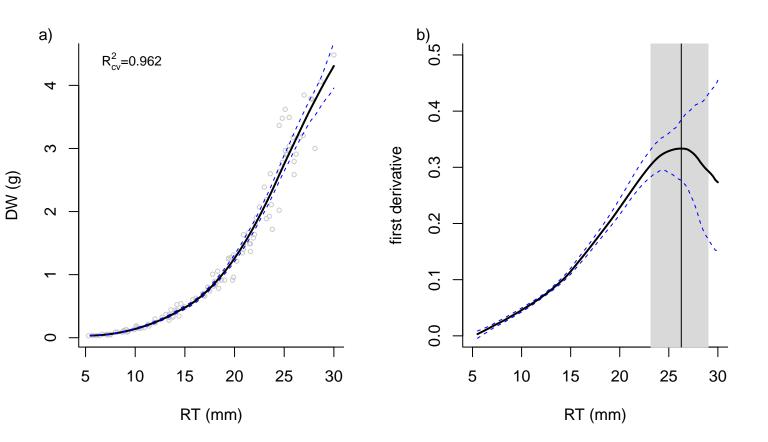
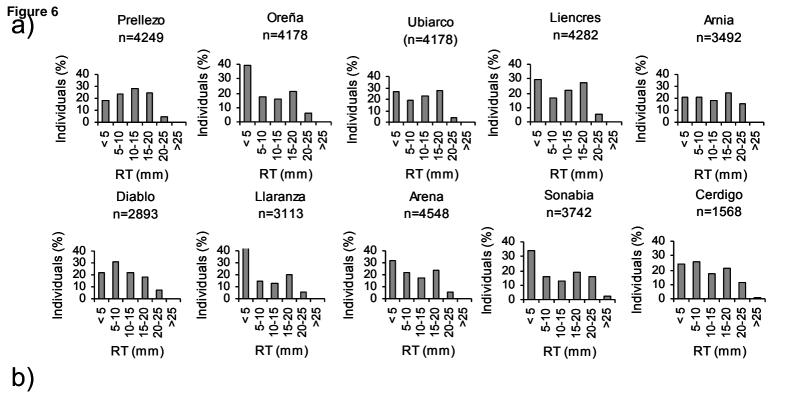
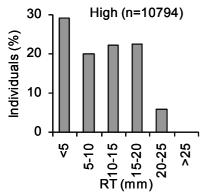
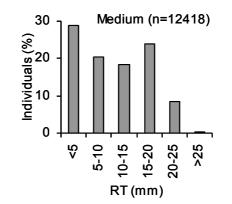


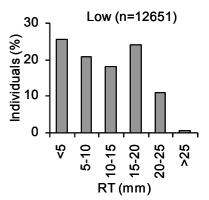
Figure 5











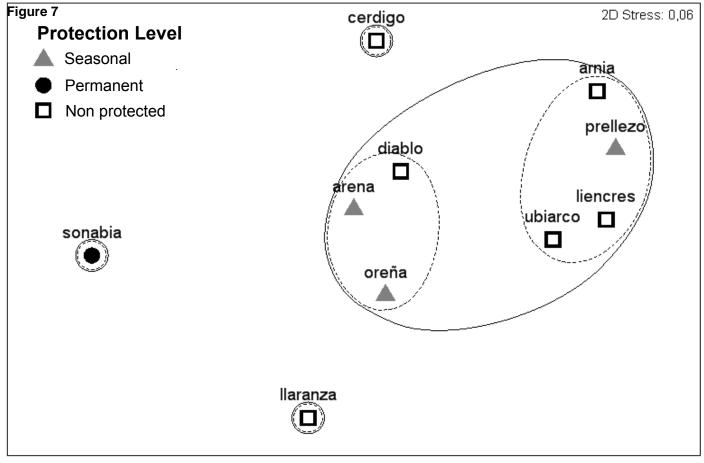


Figure A1

